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**A Simple Enquiry on Heterogeneous
Lending Rates and Lending Behaviour**

by

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A simple enquiry on heterogeneous lending rates and lending behaviour

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ABSTRACT

This paper contains an empirical analysis on the most relevant determinants of heterogeneity in Italian banks' lending rate and lending behaviour, performed with descriptive statistics and dynamic panel data analysis on the basis of a "mesoeconomic" data set specifically acquired from the Bank of Italy which contains the data on interest rates, bad debts and credit flows granted by different size categories of banks under the form of different size categories of loans in different geographic areas of the country, for the period 1990 Q1-1998 Q4. The purpose of the econometric analyses is twofold: First, investigating the possible disturbances affecting the link between policy rate and interest rates on bank credit; second, verify whether the behaviour of the largest bank loans is demand-determined. The results show that the spread between interest rate on the various classes of bank loans and the monetary policy rate is not influenced by the loan size, while it seems to be more affected by proxies for risk and competitive configuration of the banking sector. Furthermore, the behaviour of the largest size class of bank loans does show a pattern consistent with a "demand determined" behaviour.

Keywords: credit; monetary policy.

JEL Classification: E51, E52.

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A simple enquiry on heterogeneous lending rates and lending behaviour

1. Introduction

According to a well-known statement of the credit view, the monetary policy propagation process takes place both through banks' assets (loans) and banks' liabilities (money). A great deal of empirical research in that field focuses therefore on the behaviour of interest rates spreads as well as credit supply in different phases of the business cycle to different categories of borrowers. The more common empirical tools to investigate these issues are either VAR models analysing the behaviour of credit aggregates, interest rate spreads and the way these variables affect economic activity, or, more recently, dynamic panel data estimates based on individual firms (or individual banks) observations. The former directly focuses on the behaviour of macroeconomic and policy variables, the latter have the advantage of modelling individual agents and accounting for disaggregated behaviour and heterogeneity. The use of "mesoeconomic" dataset, on the contrary, might allow to analyse the interactions between some macroeconomic policy variables and the behaviour of variables associated to different categories of agents. The purpose of this paper is twofold: first, performing a simple analysis (both descriptive and econometric) on the possible disturbances affecting the link between policy rate and interest rates on bank credit. Since we are interested in heterogeneity, the second purpose of the paper aims at investigating a very relevant element of heterogeneity, i.e. verifying whether the behaviour of the largest size class of loans is demand-determined. To do that we will focus on the share of the largest size class of loans over the total: if the behaviour of this variable appears to be demand-determined, then, for a given level of credit supply, the share of the other size classes is a sort of "residual" variable, and bank credit would be allocated according to a sort of implicit hierarchy according to the size class of loans.

All the analyses are performed with a "mesoeconomic" data set acquired from the Bank of Italy, whose main characteristics are explained in detailed in the appendix. The next section briefly discusses the motivations for this methodological choice, after presenting a brief survey on the empirical research on agents' heterogeneity in the transmission mechanism of monetary policy for Italy. Section 3 and its subsections contains a preliminary and descriptive analysis of the dataset employed here. Section 4 and its subsections contain the econometric analyses. Section 5 contains a few concluding remarks.

2. Empirical research on monetary policy transmission in Italy

A common feature of the Italian empirical analyses on monetary policy transmission is the emphasis on agents' heterogeneity, credit market structure and, to a lesser extent, geographic differences

and inequalities among different regions. A great deal of this research has been performed by economists from the Research Department of the Bank of Italy and published on "Temi di Discussione", a publication from the Research Department of the Bank of Italy (for instance, Angeloni *et al.*, 1995, Baffigi *et al.*, 1999, Buttiglione *et al.*, 1994, Cottarelli *et al.*, 1997, De Bonis *et al.*, 1997, Focarelli *et al.*, 1998, Rinaldi Russo *et al.*, 1999), which, apart from being a prestigious Italian research institution, enjoys a privileged access to the generality of individual banking data for research purpose while guaranteeing, at the same time, strict confidentiality of the data, according to the existing regulations. Angeloni *et al.* (1995), by performing impulse response functions with aggregate time-series data referred to different size classes of banks show (for the sample period from 1987,1 to 1993, 12) that large-size banks tend to raise their lending rates more than small banks in periods of tight money and associate this phenomenon to monopoly power in local markets and customer relationships. Surprisingly they also find that the impact of monetary policy seems to be stronger in large companies than in the smaller ones, due, again, to the peculiar kind of customer relationships existing in local credit markets. Apart from these peculiar results, Angeloni *et al.* obtain evidence consistent with the credit view: the interbank market is weakly exogenous to the bond market, the interbank and bond market are weakly exogenous to the loan and deposit markets, and, similarly to the results by Buttiglione and Ferri (1994), the spread between bank loan and long term Government bond is influenced by the fact that lending rates tend to overshoot bond market rates. A certain influence on this kind of literature has been played by Cottarelli and Kourelis (1994) methodology in analysing how the degree of stickiness of bank rates varies significantly from country to country. Cottarelli and Kourelis also find that interest rates stickiness does not seem to be affected by the credit market structure. Cottarelli *et al.* (1997), after a detailed analysis of Italian structural data for the banking sector, apply Cottarelli and Kourelis (1994) methodology to analyze the degree of stickiness of the interest rates in various Italian provinces, by taking into consideration the possible non-stationarity of lending rates and using individual bank data. Again, like Cottarelli and Kourelis (1994), they find that Italian lending rates are indeed very sticky. Significant differences in the amount of bad debts among different geographic areas are found by De Bonis and Ferrando (1997), who also show that the lending rates in the different Italian provinces are higher the more concentrated are the credit markets, the higher the market shares of the largest banks, the more risky each single credit intermediary and the higher the bank operational costs. Focarelli and Rossi (1998) estimate with cointegration techniques a bank credit demand function (which is assumed to depend negatively on the cost of bank credit and operational cash flow, and positively on the firms' investments and on an alternative interest rate) which turns out to be very stable even in times of deep and frequent institutional changes. On the other hand, more difficult and subject to different heterogeneous influences seems to be the demand for bank credit in the various Italian region: for instance, in the North-East, real variables seem to determine a strong effect, while elasticity with

respect to the opportunity cost seems to be lower. The relevance of geographic areas is again emphasized by Finaldi Russo and Rossi (1999), who find localization in industrial districts to be a relevant and significant variable.

The focus of most empirical contributions in this research field is on agents' heterogeneity and its implications for the transmission of monetary shocks and for the process of aggregation of different observable variables. Obviously, all this does not seem to match with the "new econometrics" methodological approach, which postulates that the macroeconomic relations to be estimated have to be microfounded on the basis of a representative agent model. Since it is very hard to capture any heterogeneity within the representative agent framework, at least three theoretical attitudes seem to be possible in this regard. The first one consists of not taking too literally the "new econometrics" approach and still performing empirical analyses with aggregate time series based on theoretical assumptions only defined in aggregate terms. A good argument in this regard is provided by Blinder (1986) who claims that the use of specific properties associated to the representative agent's utility function may cause serious bias in the estimates, since "*for many goods, the primary reason for a downward sloping market demand curve may be that more people drop out of the market as the price rises, not that each individual consumer reduces his purchases*"¹. Another argument lies in the fact that the utility function of a representative agent is actually a "non-microfounded macroeconomic function": it could be theoretically obtained by integration on the basis of an aggregate consumption, and its analytical form is simply based on "*ad hoc*" assumptions².

Of course, a second (and very popular) theoretical attitude consists of only performing empirical studies with individual microdata, in order to properly capture the different and asymmetric effects that might be generated by heterogeneous agents and – at the same time – avoid any Lucas' critique objection. There is no doubt that this approach is very appropriate, but some residual (although relevant) problems might arise from sampling bias, since it is very difficult to exclude that some specific categories of individuals can be underrepresented (for instance, very small firms, in case one wishes to study the impact of monetary policy on the industrial sector of an economy). For all these reasons, we follow a third approach and perform empirical analyses with an intermediate or "mesoeconomic" level of data aggregation might provide a relevant source of information. Some goods arguments in this regard might also be provided by the recent literature on the statistical properties of aggregated and disaggregated data. Forni and Lippi (1997), for instance, show that the conventional mainstream econometric assumption on

¹ Blinder, cit., p. 76

² In addition Benartzi and Thaler (1995), Kahneman (1994), Shafir, Diamond and Tversky (1997) report a very extended experimental evidence showing that the actual behaviour of individuals choosing in conditions of uncertainty is characterized by the so-called "status quo bias" and can be modelled with a "kinked utility function", extremely different from any analytical form commonly employed to model a "nicely behaved" utility function.

the fact that only one macroeconomic source of shocks is common to heterogeneous agents is rejected by very extended empirical analyses. In addition, they show that many statistical properties (like, for instance, Granger causality) existing at the level of individual agents do not survive aggregation, while aggregated time series show statistical properties that are absent in the individual data that constitute them. All this suggests that "mesoeconomic" empirical analyses can be as informative as the more conventional microeconomic and aggregate time series analyses, by providing a different perspective to investigate the interaction between macroeconomic policy variables and heterogeneous typologies of agents.

3. A preliminary descriptive analysis

The graphics reported in the next sub-section and the descriptive statistics of the appendix allow already to identify some relevant structural phenomena. In the dataset, lenders are divided in three size classes: large and major banks, average size banks and small and minor banks. Loans are divided into 5 size classes. The smaller categories have been aggregated into the class "C1", which includes all loans smaller than 250 millions lira; class "C3" includes loans from 250 to 500 millions lira; class "C4" from 500 millions to 1 billion liras; class "C5" all loans larger than one billion liras. In most cases, the comparisons will be made among the three largest class sizes because (as explained in the appendix) the dataset contains many discontinuities in the statistical criteria of sampling and in the definitions of some relevant variables. Nevertheless, the comparison among the three largest size class is rather informative.

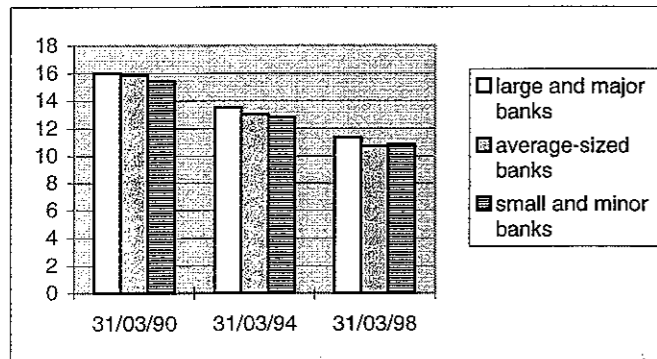
3.1 How heterogeneous are lending rates?

The first structural feature that will be considered here is the different behaviour of lending rates according to the different kinds of size classes of borrowers and geographic areas, for each given size class of borrowers. In this case, in order to make the graphics easier to read, we have compared the data referred to the first observation of each time series, the median observation and the final observation, by using histograms. This kind of graphical representation (usually not employed for data and observations of this kind, but rather to represent frequencies), although unusual, allows seeing more directly the structural differences analysed here. No precise or uniform relation can be detected between the lender size (for a given class size of loan) in the various areas, by simply observing the data and without recourse to multivariate analysis, given the potential effects of borrower risk and the demand expectations.

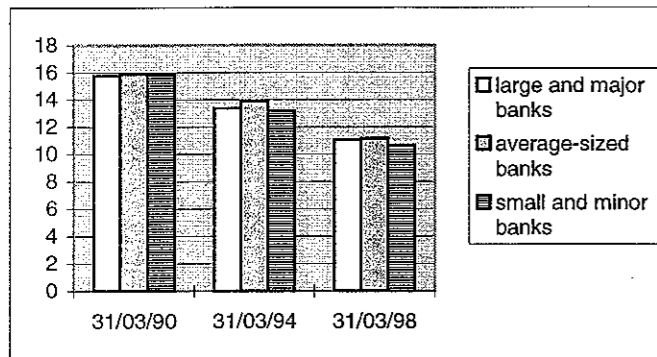
As one can see from the graphics, a higher interest rate from large and major banks is only detectable in "area 3" (corresponding to Tuscany, Marche and Umbria) for all the size classes of loans, while for the loan classes "C3" and "C4" is also detectable in "area 1" (Piedmont, Val d'Aosta and

Liguria) and in "area 5" (Abruzzi, Molise, Puglia and Basilicata). In all the other cases, there is no precise pattern (at least without using econometric methods) between the lender size and the level of the interest rates. This means that there seems to be no precise link between the lenders size and their market power. For this reason, other variables, different from the lender size are employed in the econometric analysis as proxies for the lender's market power, in particular, as shown in section 4.1, we employ the spread between maximum lending rate and minimum borrower's rate for each class size and in each geographic area.

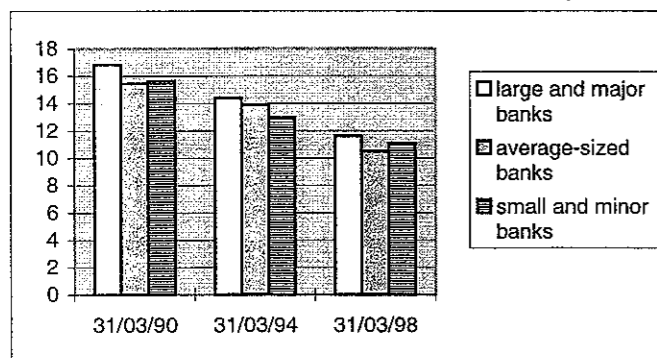
Lending rates to loan size class "c3" in area 1



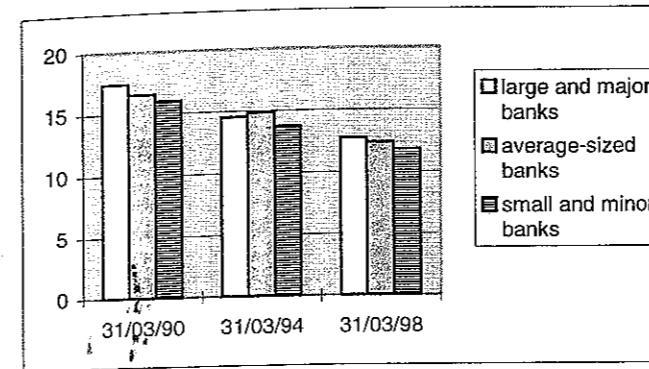
Lending rates to loan size class "c3" in area 2



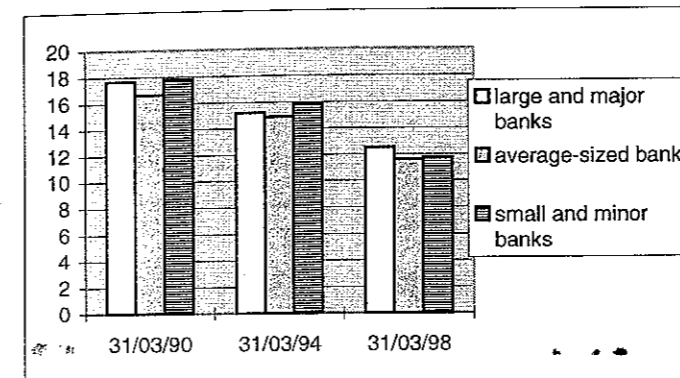
Lending rates to loan size class "c3" in area 3



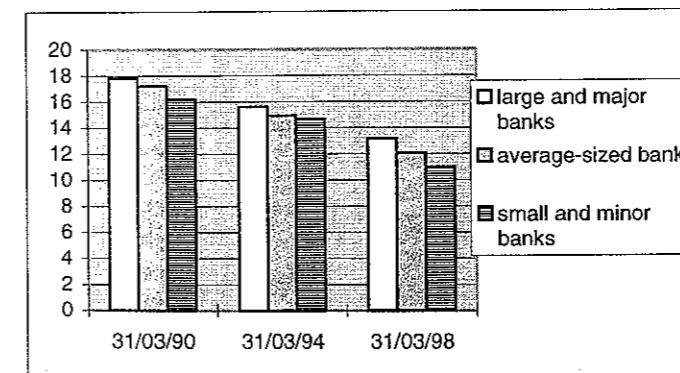
Lending rates to loan size class "c3" in area 4



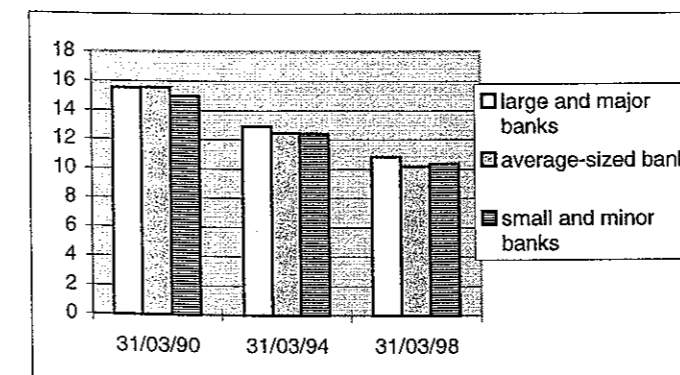
Lending rates to loan size class "c3" in area 5



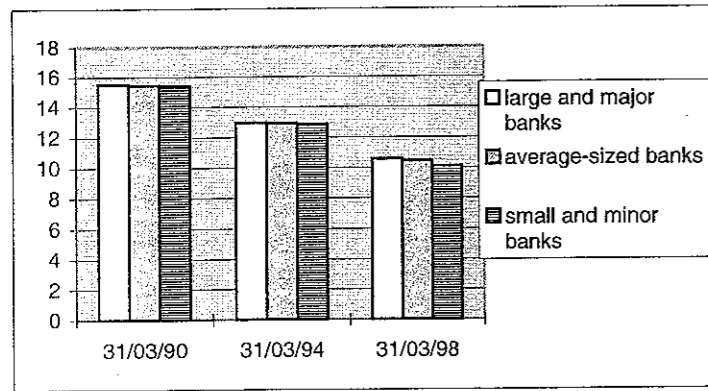
Lending rates to loan size class "c3" in area 6



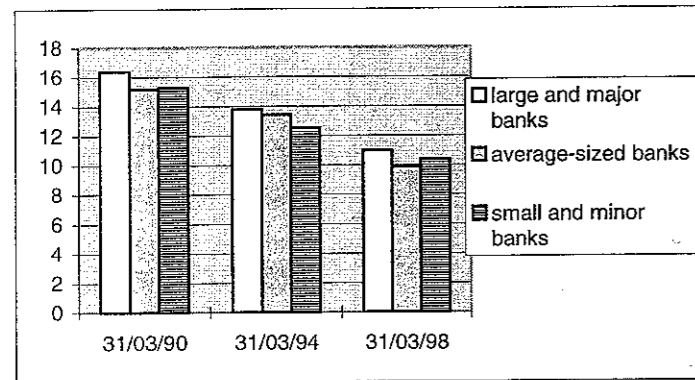
Lending rates to loan size class "c4" in area 1



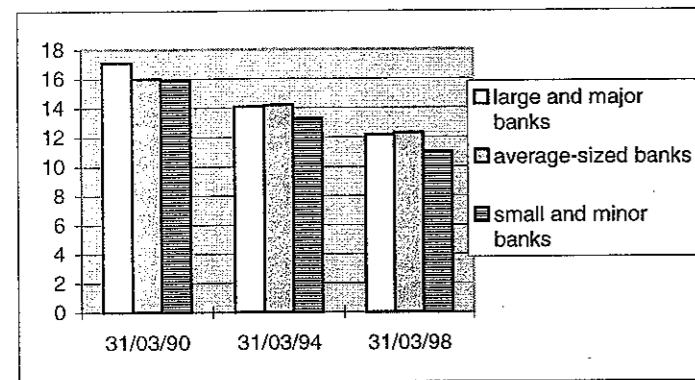
Lending rates to loan size class "c4" in area 2



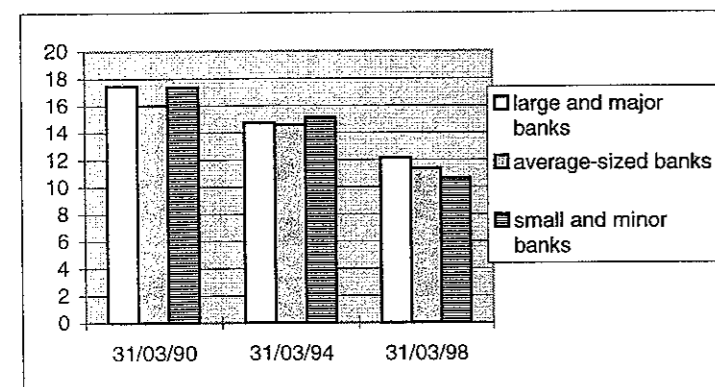
Lending rates to loan size class "c4" in area 3



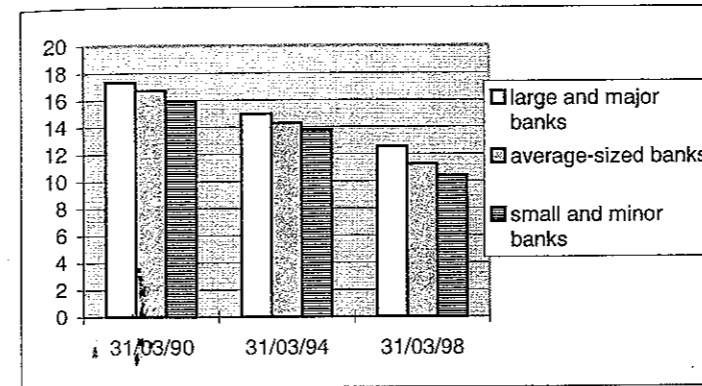
Lending rates to loan size class "c4" in area 4



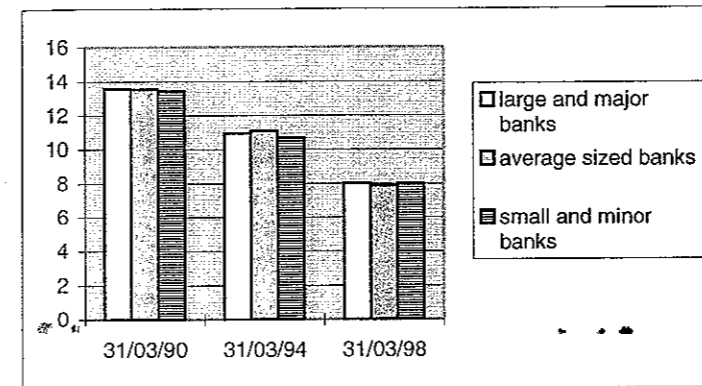
Lending rates to loan size class "c4" in area 5



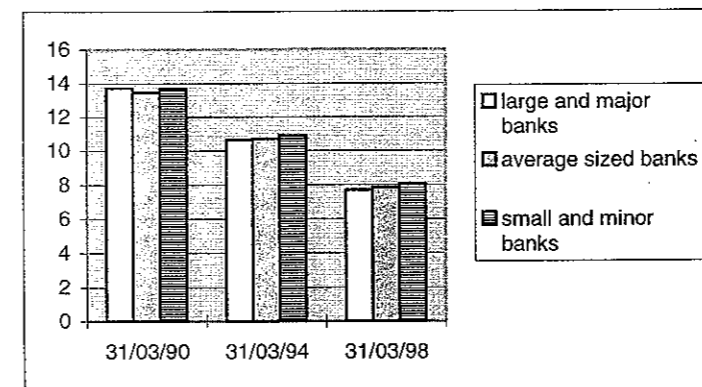
Lending rates to loan size class "c4" in area 6



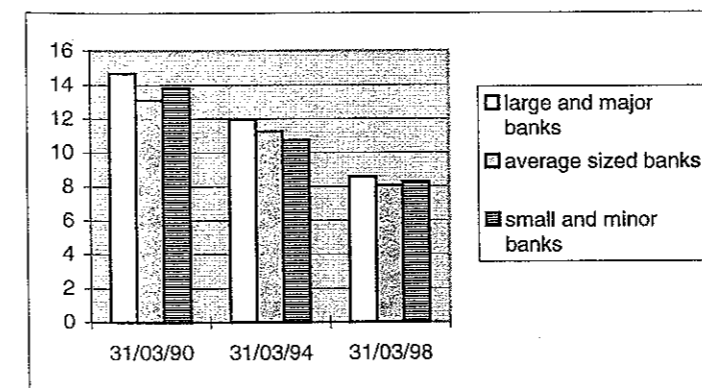
Lending rates to loan size class "c5" in area 1



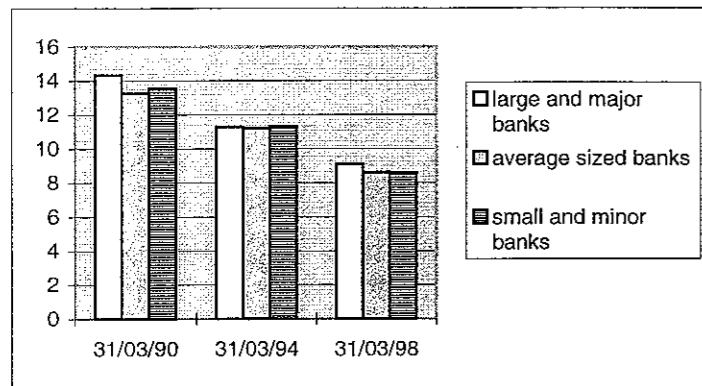
Lending rates to loan size class "c5" in area 2



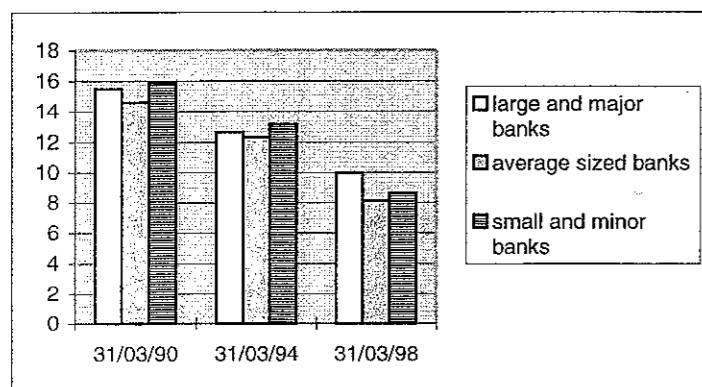
Lending rates to loan size class "c5" in area 3



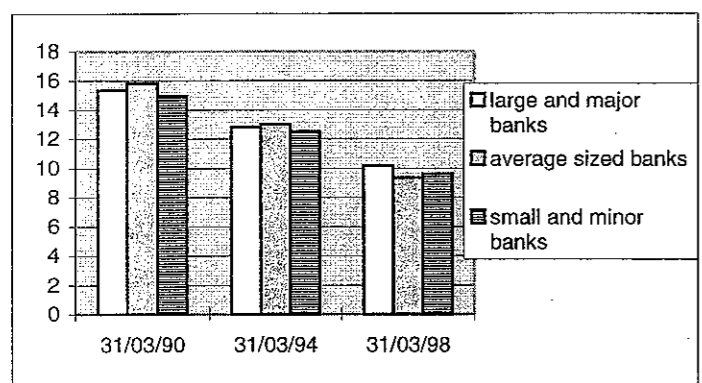
Lending rates to loan size class "c5" in area 4



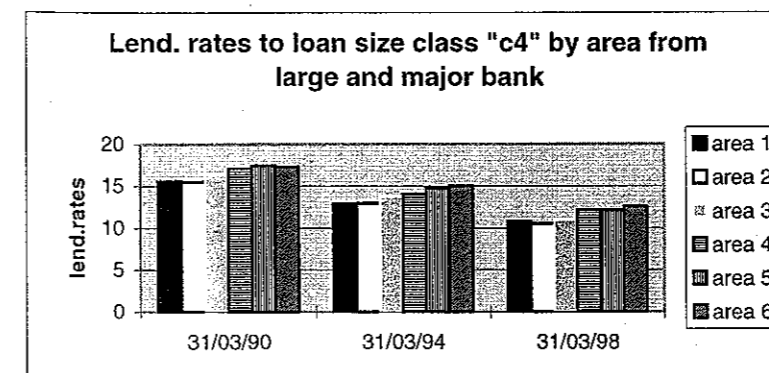
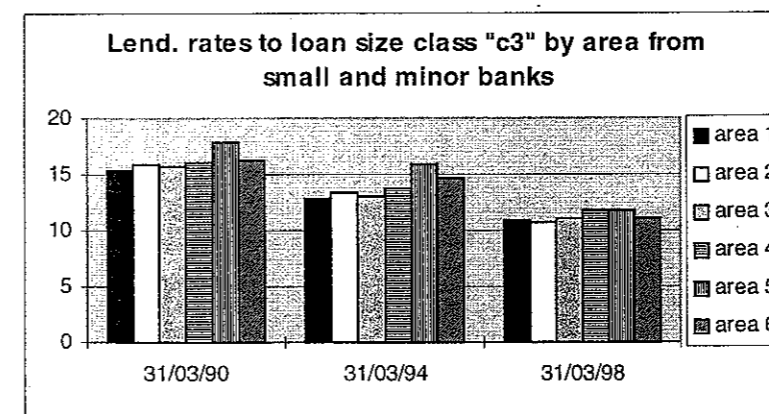
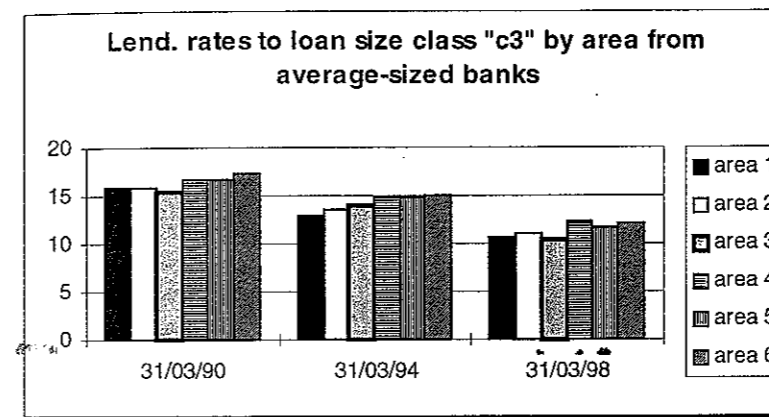
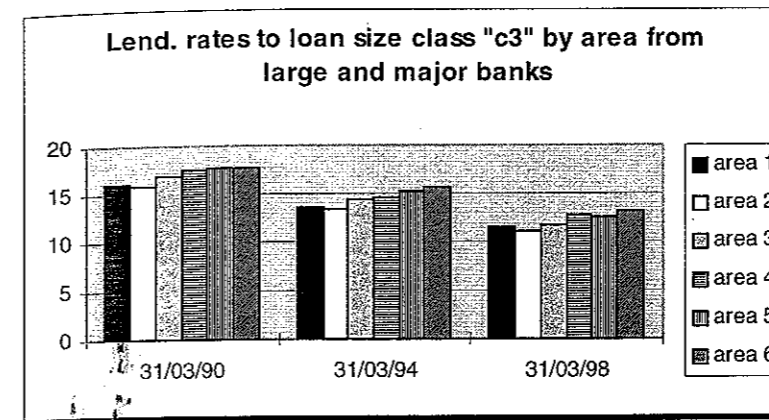
Lending rates to loan size class "c5" in area 5

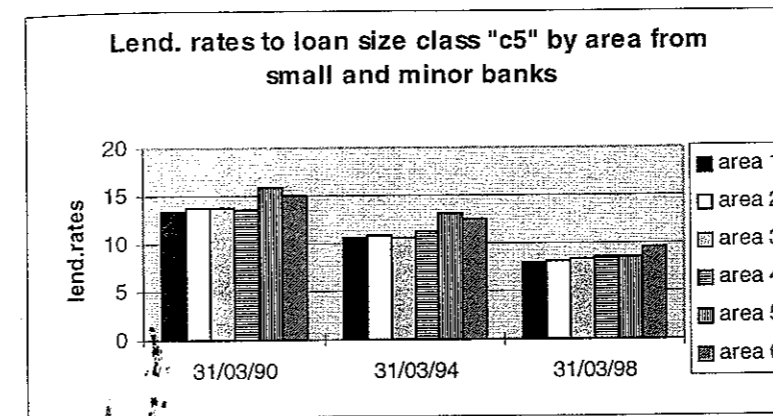
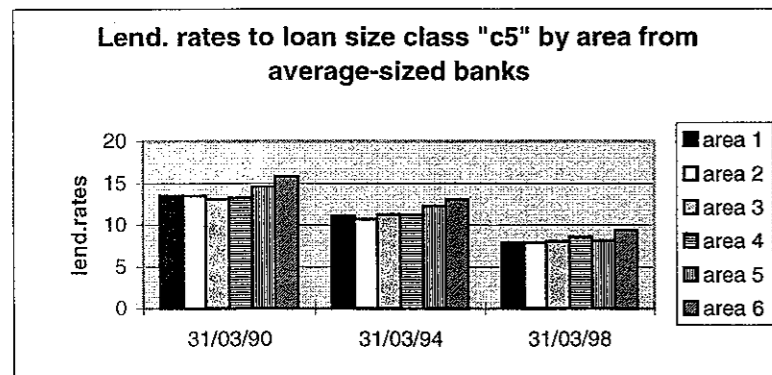
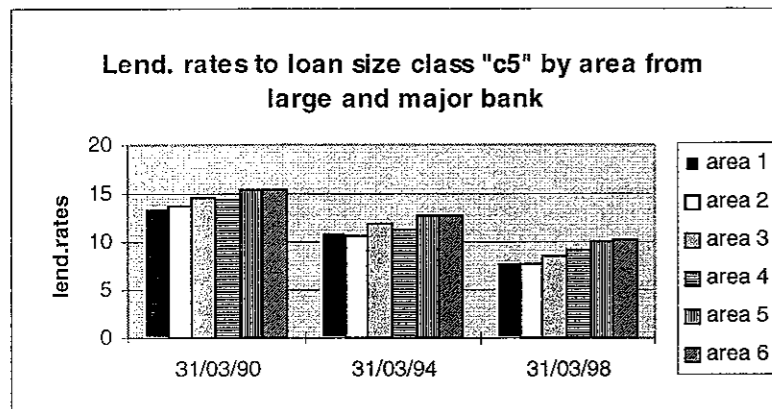
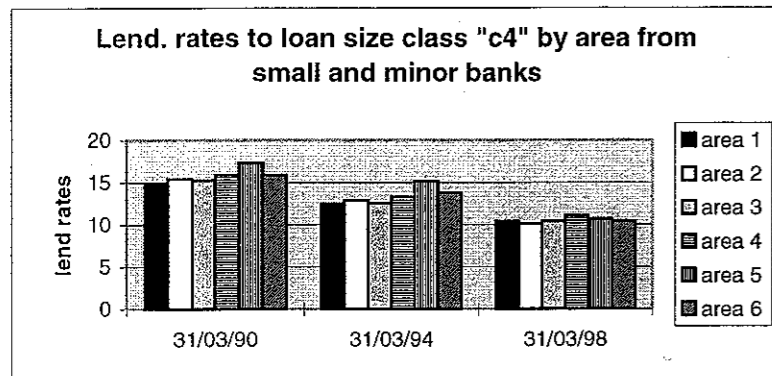
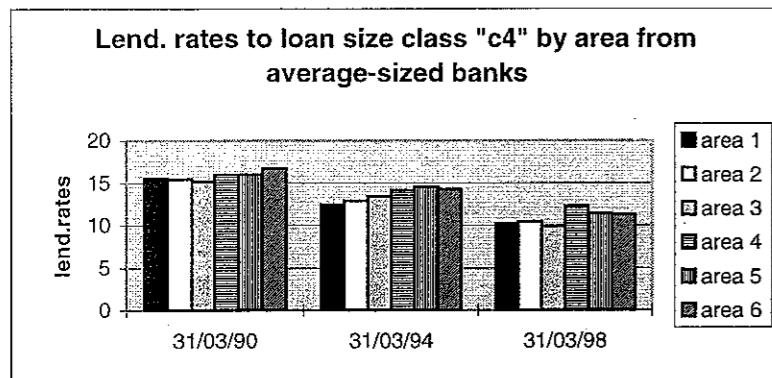


Lending rates to loan size class "c5" in area 6



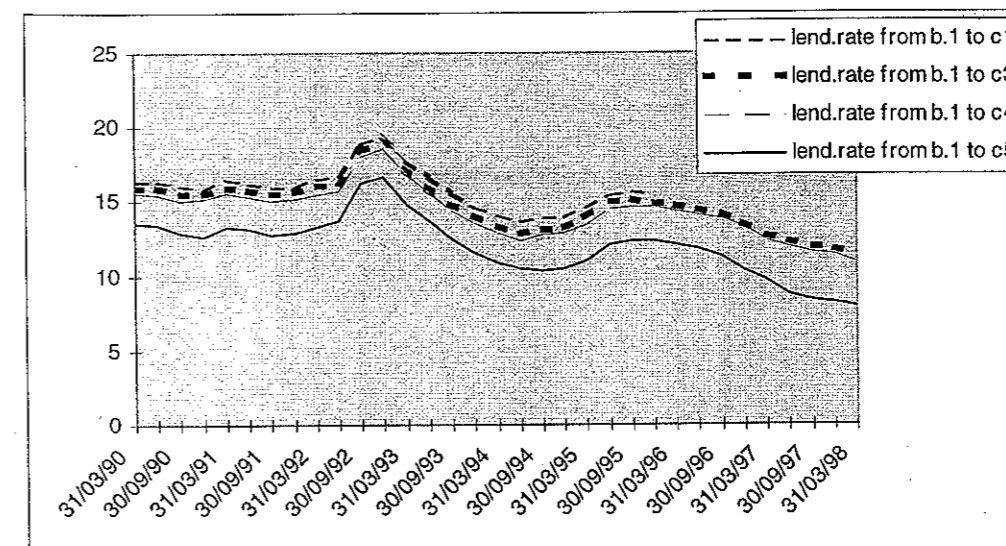
For what concerns the behaviour of interest rates for each class size of loans according to the different lenders sizes and geographic areas a certain "geographical pattern": "area 5" (Abruzzi, Molise, Puglia and Basilicata) and "area 6" (campania, Calabria, Sicilia and Sardinia) show almost constantly higher lending rates than the other areas. Also "area 4" (Lazio) would show higher lending rates than the other three areas, excepting for the loan class "C5" including the loans larger than 1 billion liras.

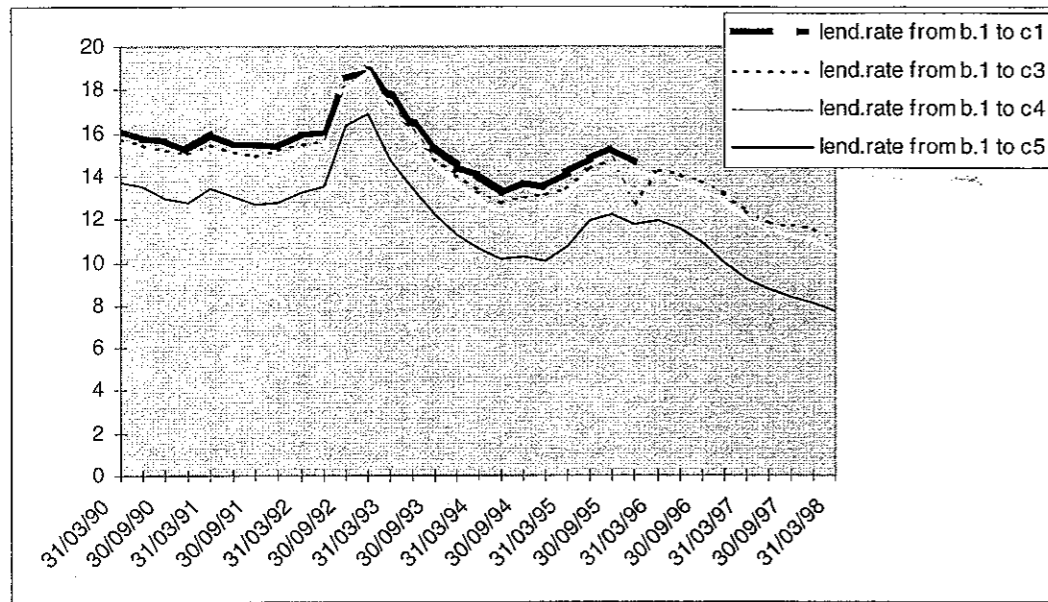




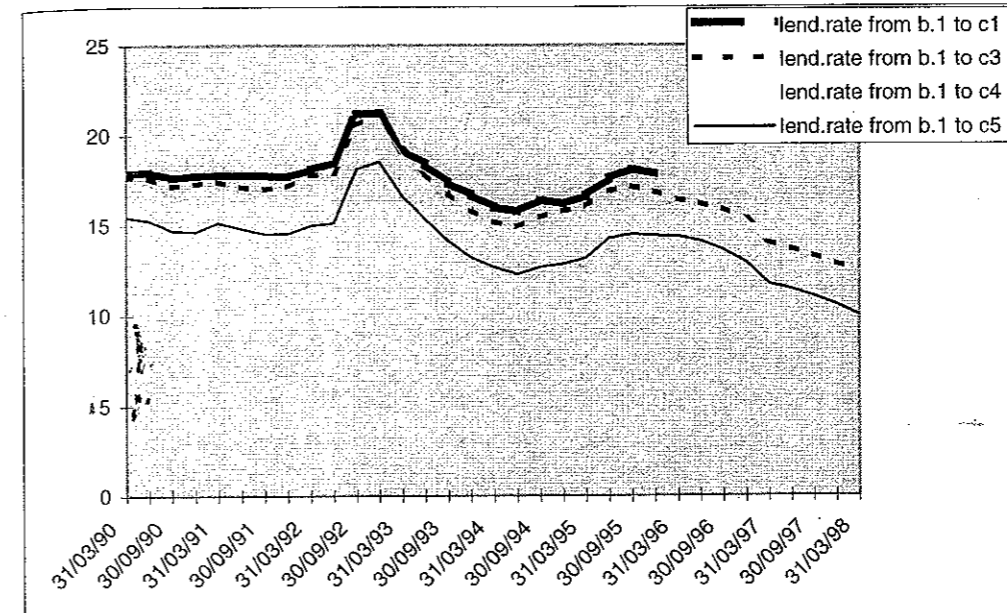
A very precise pattern, on the contrary is detectable in the behaviour of interest rates according to the different class sizes of loans in the various geographic areas (as observed for each given lender size class). In particular, the relation between the loan size class and the level of the interest rate is particularly evident if one compares the interest rate level between the loan class "C5" and the other categories. In addition, the interest rate to the smaller class size of loans is constantly higher than the interest rate to the other classes. An exception to this apparent negative relation between loan size class and lending rates is given by the lending rates from average size banks (B2) in areas 3, 5, 6 and from small and minor banks (B5) in "area 5" where the negative relation between lending rate and loan size does not exist for the intermediate size classes of loans, although it still exists between those and the two extreme loan size classes. All these data would be confirmed, and be also more evident in the descriptive statistical tables of the appendix.

Of course, the relation between loan size and level of lending rate might just be apparent and explained by different theoretical variables, such as risk and demand expectations: this is shown in the very simple econometric analysis of the next sections, where the estimates with proxies "capturing" risk and the different market power of lenders and borrowers yield completely non-significant dummies for the loan size classes, in spite of what seems to appear from the graphics.

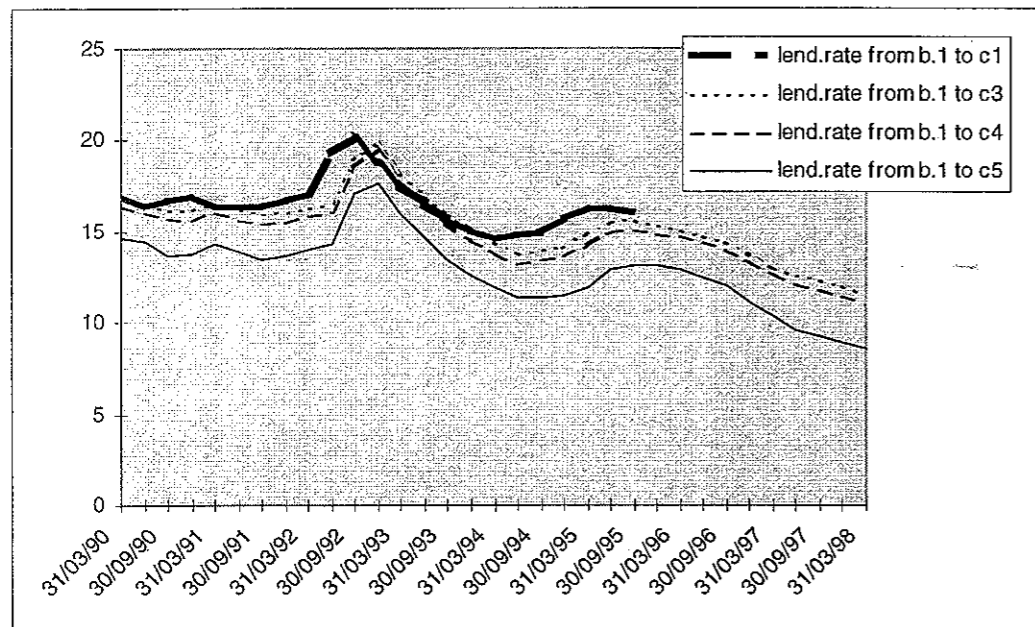




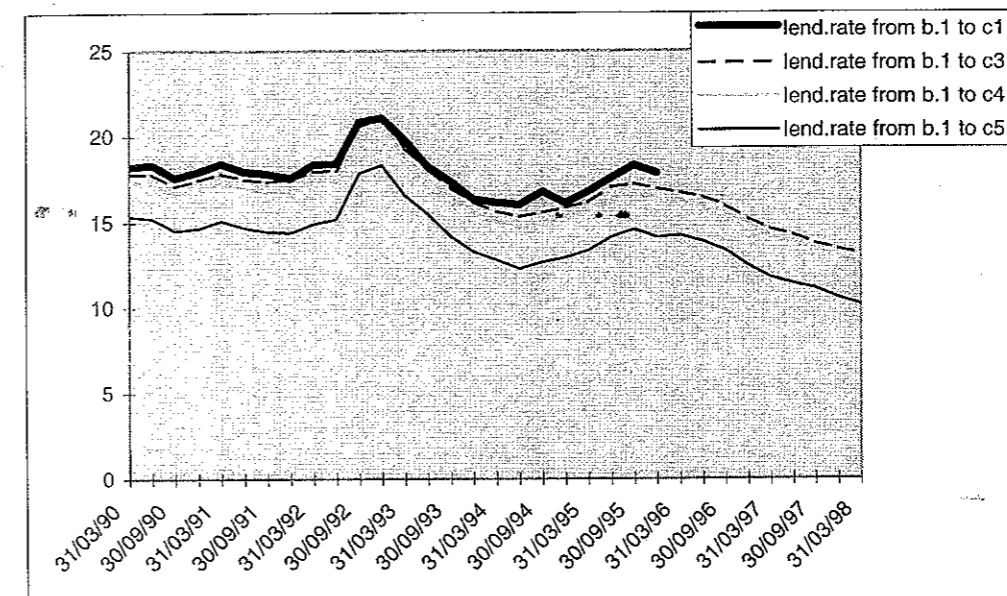
AREA 2



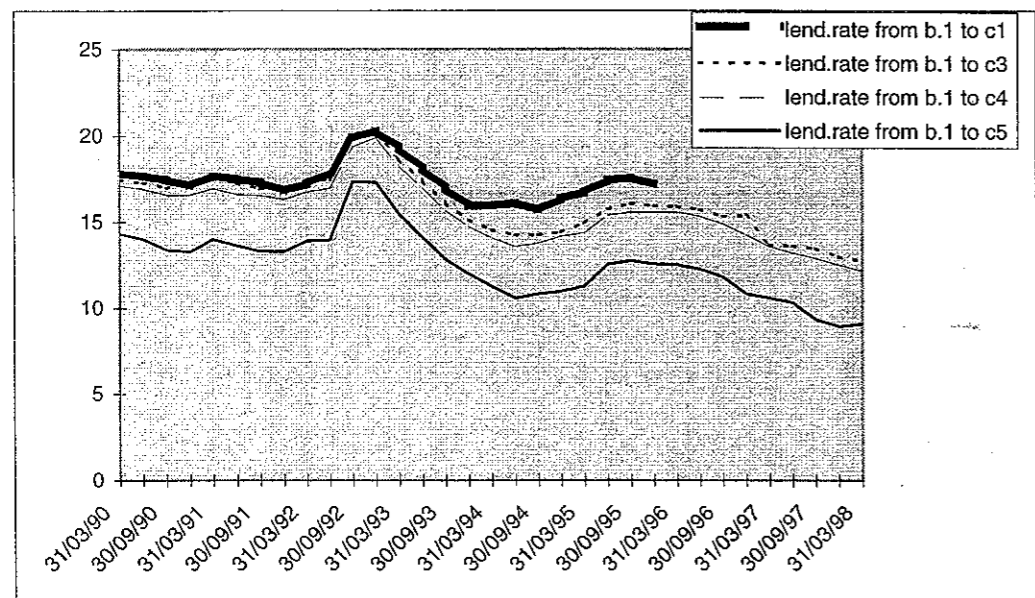
AREA 5



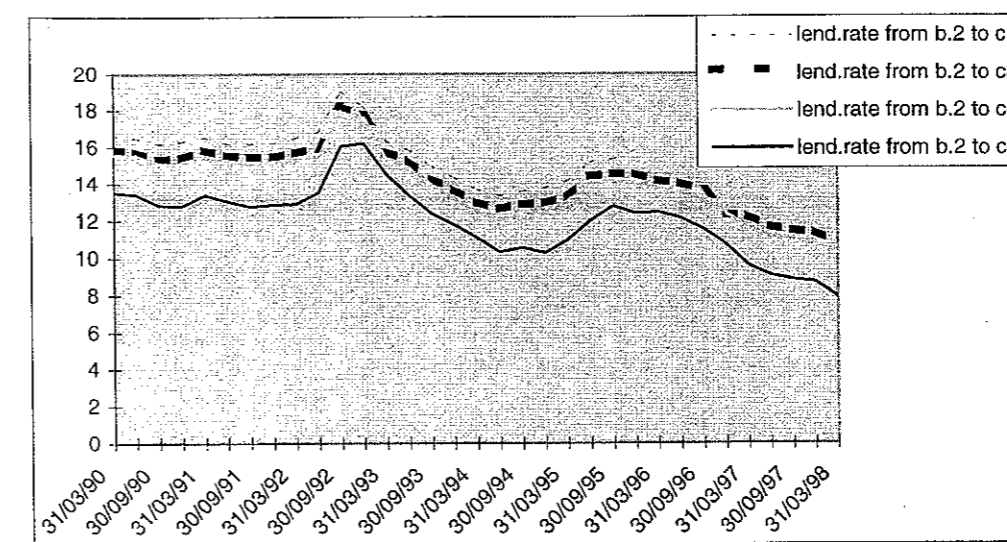
AREA 3



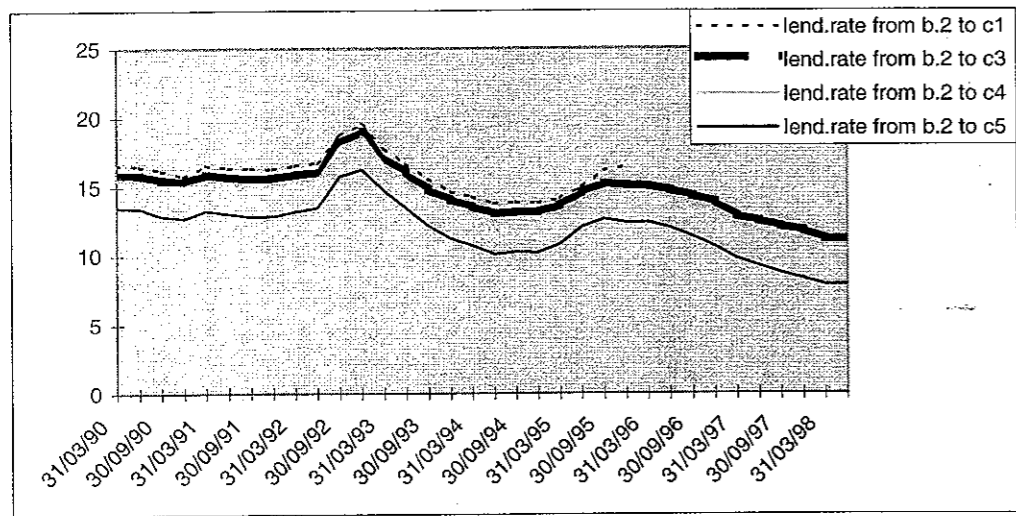
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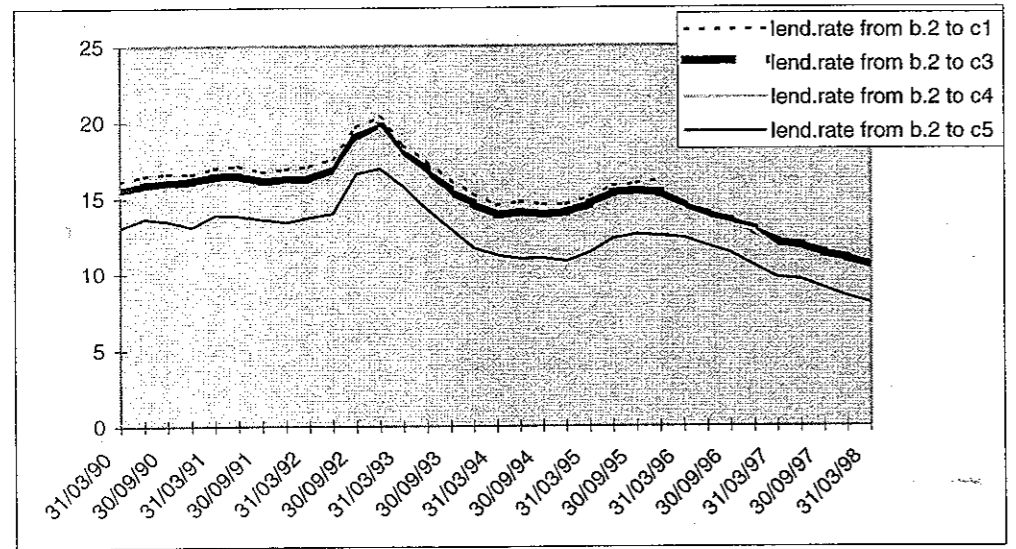
AREA 4



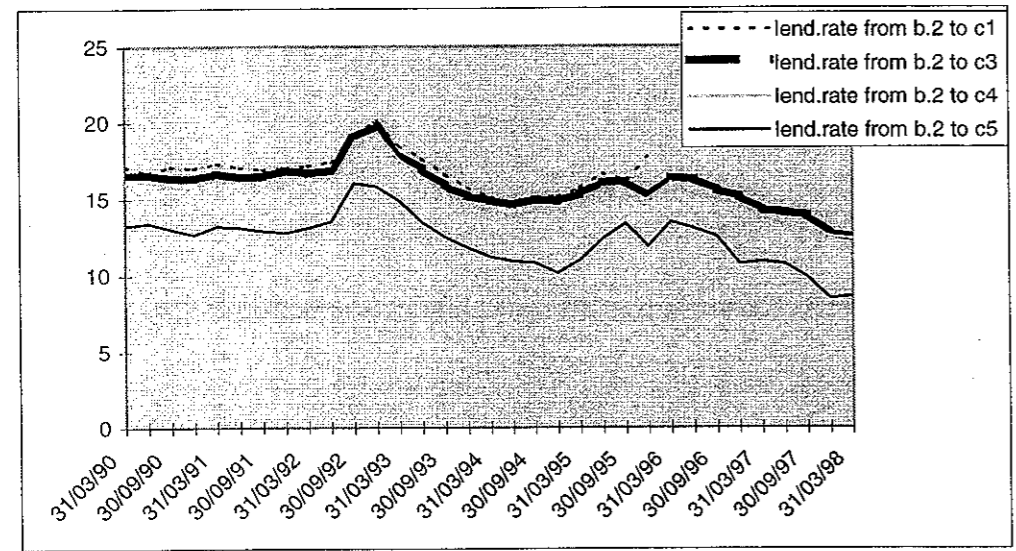
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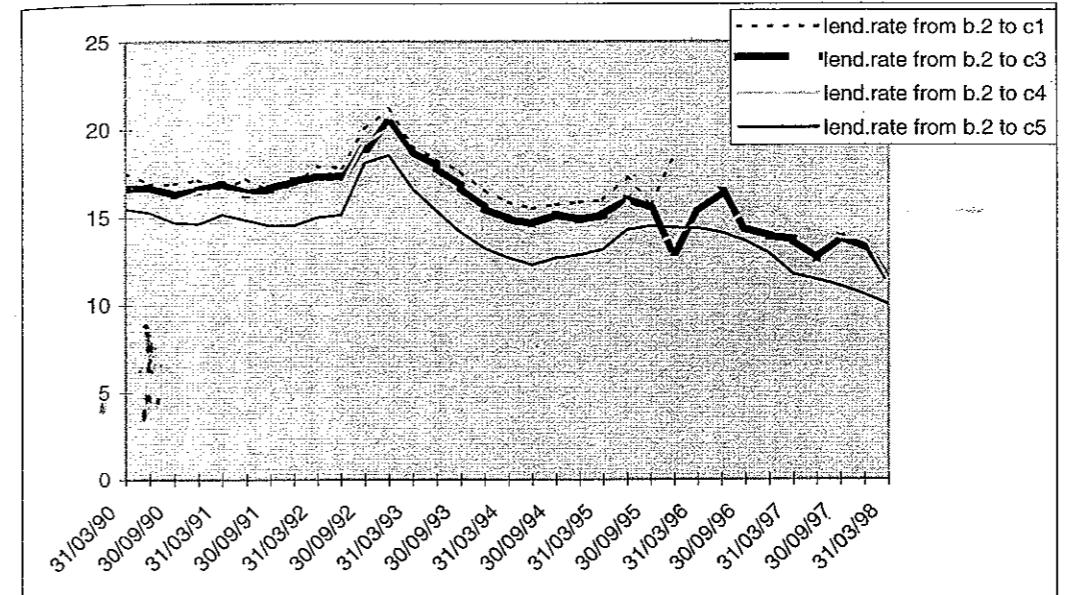
AREA 2



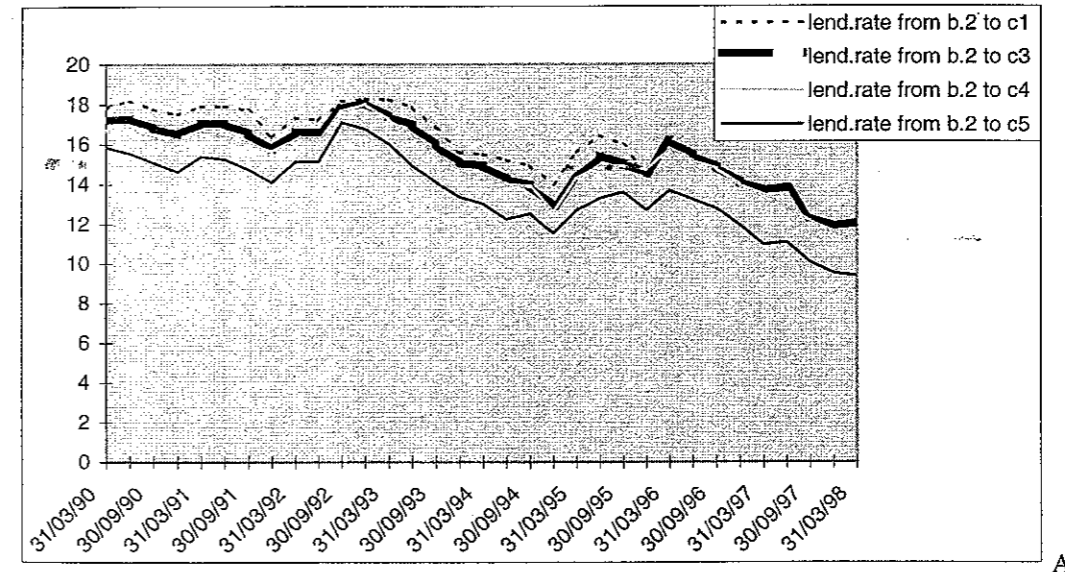
AREA 3



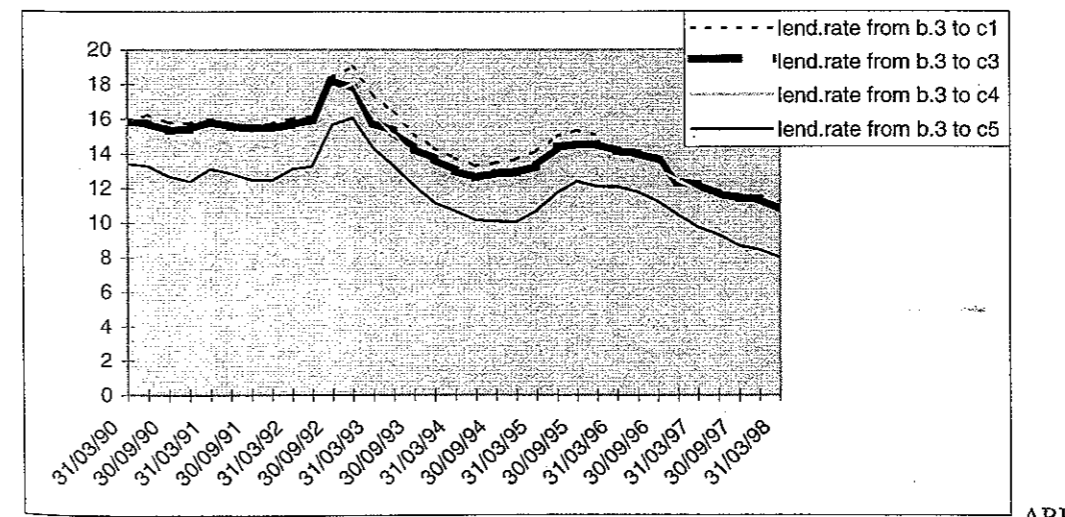
AREA 4



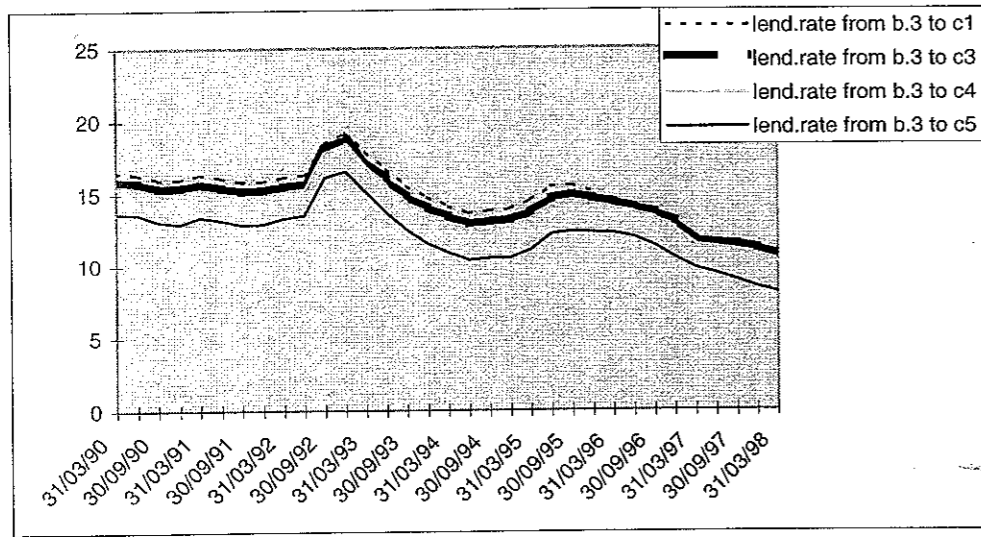
AREA 5



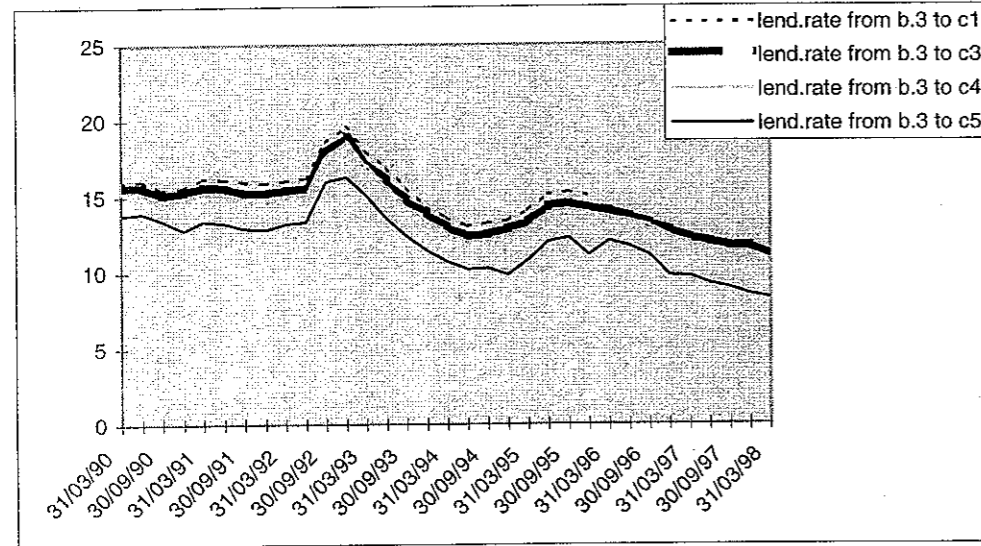
AREA 6



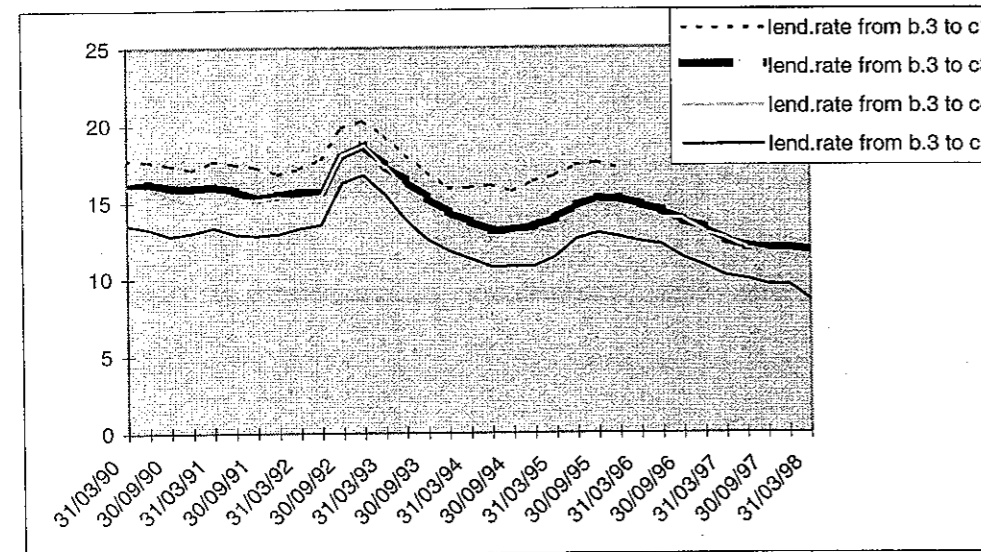
AREA 1



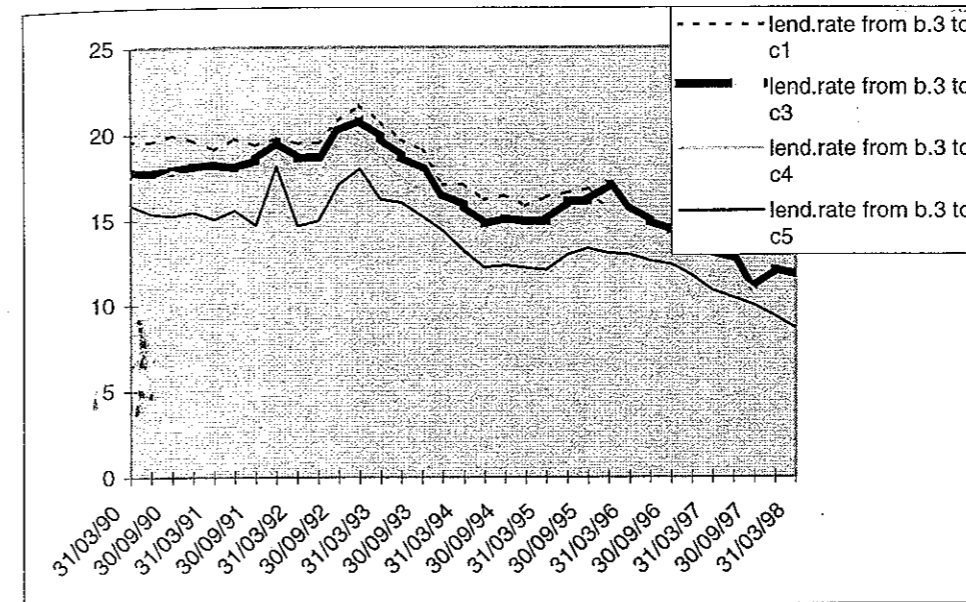
AREA 2



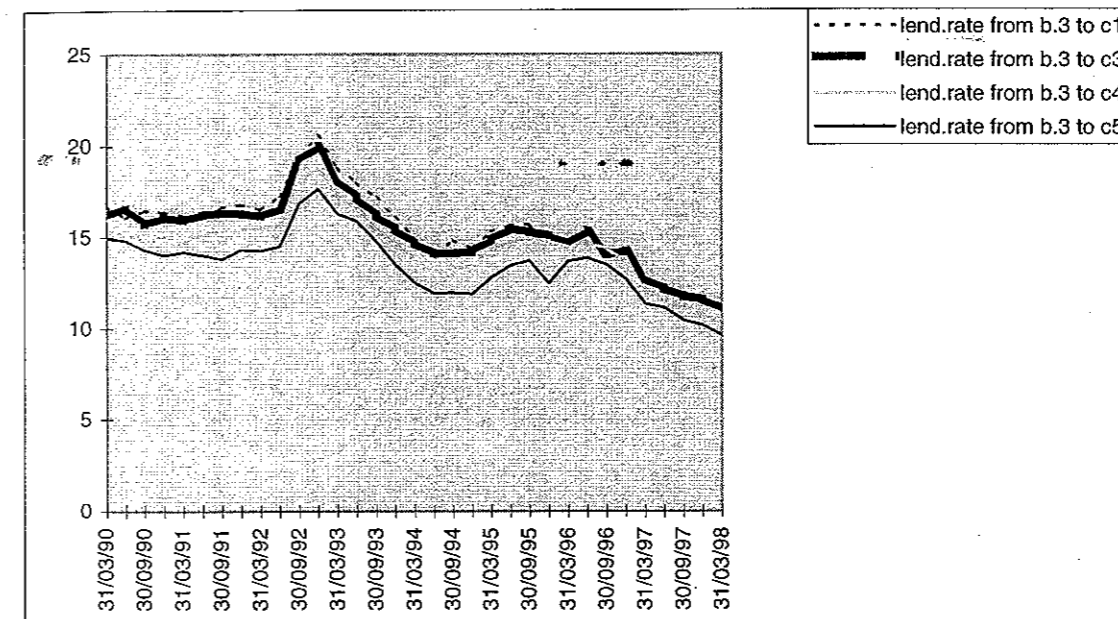
AREA 3



AREA 4



AREA 5



AREA 6

4. Some simple econometrics

Heterogeneity in bank credit and lending rates has been analysed in many different ways, as briefly discussed in section 2. The approach followed here is both unconventional and very simple. The first equation estimated (and shown in sub-section 4.1) is meant to investigate on the main determinants of heterogeneity in lending rates, for a given policy regime and is based on a simple "textbook" statement: when monetary policy is mainly based (like in the Italian context) on the control of interest rates, the rate of discount does not only display general attitude of the policy makers, but also constitutes a "topical" interest rate affecting the whole set of lending rates. The level of the latter will depend on the borrower risk, competitive context, and demand expectations. In other words, we study the spread between each specific lending rate (from each size class of lenders to each loan size class, in each geographic area and

for each quarter) and the discount rate in order to verify to what extent appropriate proxies for the category-specific risk, for the relative market power of lenders and borrowers, size dummies for lenders and loans, geographic dummies can explain heterogeneity in these categories of lenders. Most contributions in the literature briefly described in section 2 are based on *a priori* specifications or sometimes apply the "general-to-specific" methodology in a simplistic way by determining the lag structure of the model only on the basis of the t-statistics and without performing appropriate joint test for linear restrictions on a general "unrestricted" model. In this paper, we attempt to apply in a rigorous way Hendry (1985, 1988) and Harvey (1989) methodology: the final "parsimonious" specifications are strictly determined by applying joint "variable deletion tests" in a "general unrestricted" model with four time lags: The variables appearing in the estimates with lower levels of significance have not been eliminated because the zero-restrictions on their coefficient have been rejected in preliminary general "variable deletion tests". This does not strictly apply to the dummies, which have been tested jointly for the sake of our comments, in order to confront their values and level of significance. The appendix contains therefore the "general unrestricted" specifications of the models and the "variable deletion tests" performed to obtain the final specifications. Of course, by assessing the level of significance of the various regressors one has to consider that we are dealing with very volatile (and heterogeneous) financial variables. Also for this reason all the tests have been performed in the White "robust to heteroskedasticity" version (White, 1980, 1984). In a sense, the results of sub-section might be considered an extension of the descriptive analysis of section 3, since the point is not to identify a precise behaviour equation, but rather to measure and compare the impact of some theoretical and structural variables on the spread between lending rate and monetary policy rate.

The second equation estimated is meant to detect the possible existence of a hierarchy in lending behaviour and again is based on a very simple consideration. While it is certainly not possible to associate the loan size class to the size of the borrower, the largest loan size class is certainly more affected than the other two by the behaviour of large and powerful companies with higher market power. Therefore it seems reasonable to ask ourselves whether the share (and not the absolute amount) of loans granted to the largest class is demand determined. If so, for a given level of loans supplied by the banking system the share of loans to the whole of the other size classes is necessarily residually determined and cannot be demand-determined. For the sake of our estimates, we seek to verify whether the share of loans to the largest size class is negatively correlated to its price, positively correlated to the bad debts (since borrowers have obviously incentives to renegotiate them, or, at least, to seek an accommodative behaviour from the bank) and positively correlated to the demand expectations. This last point, while being consistent with a "demand-determined" lending behaviour, is not consistent with the phenomenon of "flight-to-quality".

The smallest loan size class ("c1", with less than 250 millions lira), as discussed in detail in the appendix, contains several elements of discontinuity in the time series and non-homogeneity in the definition of the relevant variables. For this reason this class could not be considered in the estimates, while, on the contrary, it has been included in the descriptive analysis of section 2. In any case, since this specific size class mainly includes credit to households and very small individual firms (like small shops), its exclusion does not affect the informational power of the second econometric analysis testing for the "demand-determined" behaviour of the largest size-class loans. The definition of the remaining loan size classes (c3, from 250 millions lira to 500 millions, c4, from 500 millions to 1 billion lira, c5, above 1 billion lira) have been "inherited" by our data set from the sampling criteria of the Bank of Italy and partly reflect (as discussed in the appendix) the constraint imposed by the norms on confidentiality of banking data.

3.1 The spread between lending rate and policy rate.

Since we cannot exclude simultaneity between some of the regressors and the dependent variables, the estimates have been performed with the method of instrumental variables. The software employed is DPD (Arellano-Bond, 1988 and further versions), the instrument employed for each variable (obviously excluding the dummies) is the lagged variable itself.

TABLE 1

VARIABLES INCLUDED IN THE ESTIMATES

CONST = intercept

I2 = dummy for the average-sized banks.

I3 = dummy for the "small and minors" banks (according to the Bank of Italy classification).

c3 = dummy for the loan size class "c3" (i.e. from 250 millions lira to 500 millions lira).

c4 = dummy for the loan size class "c4" (i.e. from 500 millions lira to 1 billion lira).

c5_t = ratio between the bank overdrafts granted to the category "c5" and the total sum of bank overdrafts granted to categories "c3", "c4" and "c5".

isco = expected demand for the next 3-4 months (ISCO data, based on entrepreneurs interviews). It is a proxy for the firms' forward-looking expectations and can assume positive or negative values, since it expresses expectations on the possible increase or decrease in demand.

so_u = ratio between bad debts and used credit, for each geographic area, for each year, from each size category of banks to each size category of loan facilities; because of the discontinuity in the criteria of statistical sampling followed by the Bank of Italy, (briefly discussed in the appendix) the ratio does not include the loan credit commitment. The variable has been used as a proxy for the "observation-specific" risk.

ta_d = tat - tan

tan = interest rate on loans by the Bank of Italy, inclusive of any additional charge.

tat = interest rate on loan facilities, for each geographic area, for each year, from each class size of banks to each size class of loan facilities.

tdif = tma - tmi

tma = maximum observable alternative interest rate for the loans facilities of each size category of banks for all the size categories of overdrafts, in each geographic area in each year.

tmi = minimum observable alternative interest rate for the customers demanding loans facilities of a given size from all the possible alternative size categories of banks in each geographic area, in each year.

Each observation unit is the credit (and interest rates) from each size class of lenders to each loans size class, in each geographic area, for each quarter. There are 3 size classes of lenders 3 loan size classes, 6 geographic areas, 33 quarters, which makes it 54 longitudinal units observed for 33 periods. The variable "so_u" is a proxy for the risk specific to each specific class of borrower from each size class of lender, in each geographic area in a given year and quarter. Therefore its level (and possibly its increase) expected to be positively correlated with the spread "ta_d", since the lending rate must include a risk premium. The variable "tdif" is a proxy for the competitive context. A higher spread between the maximum lending rate available to the relevant size class of lenders and the minimum lending rate available to the relevant loan size class reflects a situation of higher competition among banks and larger possibilities of choice for the borrowers. The rationale for using this variable instead of only using more conventional "structural" variables or size variables to capture the market power of the counterparts lies in the fact that the competitive context might change much more quickly than structural variables, sometimes under the effect of strategies performed by medium-sized and even small firms and banks. This variable is then expected to be negatively correlated with the spread "ta_d". The variable "isco" represent the general expectations of the economy: if one accepted the assumption of rational expectations it should be common to all the agents (since, in particular, this variable is based on publically available interviews on entrepreneurs). It constitutes again an element of general risk evaluation, at a macroeconomic level, however it turns out to be totally non significant in the equation for "ta_d". On the basis of all the considerations made before, the following general unrestricted model has been estimated:

$$\begin{aligned}
 ta_d_{ijt} = & CONST + \sum_{k=1}^4 \alpha_{t-k} ta_d_{ijt-k} + \sum_{k=0}^4 \beta_{t-k} tdif_{ijt-k} + \sum_{k=0}^4 \gamma_{t-k} so_u_{ijt-k} + \\
 & + \sum_{k=1}^4 \phi_{t-k} isco_{t-k} + \\
 & + \text{dummies for the loan sizes} + \\
 & + \text{dummies for banks size} + \\
 & + \text{dummies for the geographic area} + \\
 & + \text{white noise}
 \end{aligned}
 \tag{1}$$

The results of the estimates and the "variable deletion tests" which have determined the final and "parsimonious" specification are reported in the appendix. The final "parsimonious" specification yields the estimates of Table 2.

TABLE 2
FINAL "PARSIMONIOUS" SPECIFICATION
EQUATION FOR THE SPREAD BETWEEN INTEREST RATES ON BANK OVERDRAFT AND
MONETARY POLICY RATE (INTEREST RATE ON LOANS BY THE BANK OF ITALY)
The data employed for the estimates cover the period 1990 Q1 - 1998 QIV

D.P.D. RESULTS
LEVELS IV
Number of firms: 54 Sample period is 1991 QII to 1998 QI
Observations: 1512 Degrees of freedom: 1497
Dependent variable is: ta_d
Instruments used are:
CONST ta-d(-2) so_u(-1) so_u(-2) tdif(-4) tdif(-5) c3 c4
a1 a3 a4 a5 a6 IND DUMS

ONE-STEP ESTIMATES WITH ROBUST TEST STATISTICS

Wald test of joint significance:	2230.202567	df = 12		
Wald test - jt sig of ind dums:	1.571201	df = 2		
Var	Coef	Std. Error	T-Stat	P-Value
CONST	1.302663	1.159927	1.123056	0.261414
ta_d(-1)	0.748006	0.124735	5.996774	0.000000
so_u	5.543778	6.650326	0.833610	0.404501
so_u(-1)	-5.534232	6.553838	-0.844426	0.398431
tdif(-3)	-0.958507	1.332951	-0.719086	0.472088
tdif(-4)	0.750205	0.976935	0.767917	0.442537
c3	0.057856	1.302231	0.044429	0.964563
c4	0.072287	1.034254	0.069893	0.944279
a1	-0.032614	0.085680	-0.380652	0.703462
a3	0.125009	0.094511	1.322695	0.185937
a4	0.249281	0.109245	2.281853	0.022498
a5	0.393597	0.210871	1.866527	0.061968
a6	0.303455	0.241046	1.258909	0.208063
I2	-0.247605	0.198526	-1.247216	0.212318
I3	-0.392050	0.325261	-1.205340	0.228072

Robust test for first-order serial correlation: -2.878 [54]
Robust test for second-order serial correlation: -0.386 [54]

The variables referred to the overdue bank credits ("so_u" and "so_u(t-1)") are only significant at a level of confidence of around 60%, while "tdif(-3)" and "tdif(-4)" (capturing the bargaining power of the banks) are only significant at a level of confidence of 53% and 56% respectively, although they were much more significant in the "general unrestricted model", which also did not support a linear restriction

setting their coefficient equal to zero, in spite of their low level of significance shown by the t-statistics. Therefore, their low level of significance might be due to the contemporaneous presence of other dummies (which are also meant to "capture" other relevant effects associated to the specific behaviour of lenders in their approach to risk or investment strategies and the specific risk of some geographic areas). By looking at the coefficients of $tdif(-3)$ and $tdif(-4)$ one could interpret the dependent variable "ta_d" as negatively affected by the variable "tdif(-3)" (with a coefficient close to -0.20) and also negatively affected by the difference "tdif(-3)-tdif(-4)" (with a coefficient of close to -0.75), although these two variables are not very significant. The variable referred to the size dummies of the bank loans are not significant at all, although they were significant in the general unrestricted model with a level of confidence of 76% and 85%. However, since for the sake of our comments all the dummies have been considered jointly, the loan size dummies appear in the estimates. The results suggest that the persistently higher interest rate set by the banks on the smaller sized overdrafts might not be simply associated to the size of the credit contract, but could reflect the impact of other more conventional variables, associated to the riskiness of the borrower, the competitive context (captured by the variable "tdif"), the size of the lender, whose dummies turn out to be significant at a level of confidence of 77% and 79%, and, partly, to the geographic area (specially for what concerns the region Lazio, whose dummy is significant with the level of confidence of 97%, but also for the areas 3, 5 and 6, whose dummies are significant at the levels of confidence of 82%, 94% and 80% respectively). Finally, the final parsimonious specification marginally shows some marginal first-order residual correlation.

3.2 Are largest loans demand-determined?

As explained before, our conjecture is that the *share* of largest loans over the total amount of loans granted to the sum of the size classes "c3", "c4", "c5" is "demand determined" if it is negatively correlated with its own interest rate, positively correlated with the demand expectations, and reflect an accommodative behaviour from the banks. On the basis of the previous considerations, the following general unrestricted model has been estimated (see the appendix for the details and the variable deletion tests).

$$c5_t_{ijt} = \text{CONST} + \sum_{k=1}^4 \eta_{t-k} c5_t_{ijt-k} + \sum_{k=0}^4 \theta_{t-k} \text{tat}_{t-k} + \sum_{k=0}^4 \lambda_{t-k} \text{tan}_{ijt} + \sum_{k=0}^4 \sigma_{t-k} \text{so_u} + \sum_{k=0}^4 \xi_{t-k} \text{tma} +$$

+ dummies for banks size +

+ dummies for the geographic area +

+ white noise (2)

As an alternative lending rate "tma" and not "tmi" has been employed: this constitute a

conservative choice to test for "demand determined" behaviour of the dependent variable, since "tma" reflects the incentive of the banks rather than borrowers. By looking at the final "parsimonious" specification reported below, we see that the own interest rate with three lags "tat(-3)" is largely significant and has a negative coefficient, as expected. By looking at the lag structure of the variable "so_u" and at the values of its coefficient one could say that the share of loans to class c5 is roughly positively correlate with an increase of the bad debts over the last 2-3 quarters. However the second and third lag are significant at the level of confidence of 88% and 83% respectively, while the 4th lag is not very significant (although in preliminary a variable deletion test the zero-restriction on its coefficient was rejected). Nevertheless, by looking at the (very close in absolute value) coefficients of "so_u(-2)" and "so_u(-3)", what seems to be significant is rather the "increase" in the bad debts: this again is consistent with an accommodative behaviour of banks and, therefore, a demand determined share of "c5". Also the positive (and significant at the level of confidence of 83%) if the variable "isco" is consistent with a "demand determined" behaviour. None of the other variables are significant and it is interesting to remark in this case that no dummy variable turned out to be significant.

TABLE 3
FINAL "PARSIMONIOUS" SPECIFICATION
EQUATION FOR THE RATIO BETWEEN THE LARGEST SIZE CLASS OF LOAN CREDITS
AND THE TOTAL CREDIT GRANTED FOR ALL SIZE CLASSES
The data employed for the estimates cover the period 1990 QI - 1998 QIV

D.P.D. RESULTS

LEVELS IV

Number of firms: 18 Sample period is 1991 QII to 1998 QI
Observations: 504 Degrees of freedom: 497

Dependent variable is: c5_t

Instruments used are:
CONST c5_t(-2) tat(-4) so_u(-3) so_u(-4) so_u(-5) isco(-1)

ONE-STEP ESTIMATES WITH ROBUST TEST STATISTICS

Wald test of joint significance: 31994.696540 df = 6

Var	Coef	Std. Error	T-Stat	P-Value
CONST	0.026569	0.009274	2.864828	0.004172
c5_t(-1)	0.977520	0.009422	103.745459	0.000000
tat(-3)	-0.000630	0.000262	-2.403433	0.016242
so_u(-2)	0.113115	0.072979	1.549970	0.121149
so_u(-3)	-0.103423	0.075927	-1.362147	0.173151
so_u(-4)	-0.011530	0.021300	-0.541309	0.588295
isco	0.000100	0.000073	1.373599	0.169566

Robust test for first-order serial correlation: -0.154 [18]
Robust test for second-order serial correlation: 1.412 [18]

4. Concluding remarks

This paper contains an empirical investigation on the various potential sources of heterogeneity in Italian bank lending rates, according to the lender size, the size class of bank loans, and the geographic area, for a period of observation going from 1990Q1 to 1998Q4, using a data set specifically acquired from the Bank of Italy.

The first econometric analysis investigates the behaviour of the spread between the lending rate (from each size class of lenders to each size class of loans in each of the six areas, in the estimating period) and the monetary policy rate, since in this context, the amount of the spread might be interpreted as a diversified and heterogeneous reaction to a given "policy measure", such as indicated by monetary policy interest rate. While the preliminary descriptive statistical analysis seems to show a clear and persistent negative correlation between the class sizes of bank loans and the level of lending rates, the econometric analysis shows that this correlation can be well explained by more conventional theoretical variables, such as the borrower-specific risk and a proxy for the relative bargaining power between lender and borrower. Like in other Italian contributions on this kind of literature, it is also found, for some geographic areas of the country, a persistently higher level of lending rates, while the significance of dummies for the lenders size is more dubious and the dummies for the loan size classes are certainly not significant.

The last empirical analysis finally identifies an interesting source of heterogeneity in the fact that the share of loans to the largest size group acts consistently with a "demand-determined" behaviour.

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APPENDIX

Sources and criteria for statistical surveys followed by the Bank of Italy

A) SOURCES

A detailed description of the classification criteria can be found in "Supplementi al Bollettino Statistico", Banca d'Italia, Anno V Numero 32 - 16 June 1995. The dataset employed for this paper has been provided by the Statistical Bureau of the Bank of Italy, on the basis of a specific request and application. Therefore, although it has been structured and organized on the basis of our research plan, it is necessarily based on the criteria of sampling and aggregation employed by the Bank of Italy for what concerns the definitions of the different size classes of loans from different size classes of banks. In addition, the degree of disaggregation is conditioned by the Italian regulation on banking secrecy, stating that each disaggregated observation of credit flows from particular size classes of lenders, to particular size classes of borrowers, for given geographic areas must be based at least on three observations. The same applies to the data concerning interest rates, which are recorded by the Bank of Italy on the basis of a non-compulsory quarterly report produced by a sample of 80 banks in all the country. For these reasons, only a limited degree of disaggregation has been possible. The complete dataset employed for this empirical analysis is now available on the internet at the site

www.economia.unimo.it/worppap/RePEc/mod/modena/

The data on interbank and foreign deposits and securities held by banks in their portfolios are based on reports from each individual bank to the Board of Inspection of the Bank of Italy. The data on loan credit commitments, total used credit, total granted credit and bad debts are based on notice from each individual bank to "Centrale dei Rischi" (the Department of Bank of Italy charged with monitoring the system risk). The amounts are expressed in million lira and include both the operations in liras and in foreign currency.

The interest rates on loan credits and on deposits have been obtained from the respective quarterly statistical survey, which includes observations approximately 70 banks. The data on interest rates on bank lending and deposits are only referred to the operations in liras.

B) LIMITS FOR THE STATISTICAL SURVEYS

From January 1996 the obligation to signal risks to "Centrale dei Rischi" only concerns loan cash credits and credit commitments for amounts greater than 150 millions lira, while before that date the limit was 80 millions lira. This creates a relevant discontinuity in the time series referred to the smallest loan size, which, as a consequence, could not be employed for the econometric estimates and has been included only from 1990 until 1996 in the graphics and descriptive statistics.

For the bad debts reported to Centrale dei rischi, there is no lower limit from february 1991, while the limit was 10 millions lira before 1991. This creates yet another source of discontinuity in the time series referred to the smallest loan size.

For what concerns the interest rates on loan cash credits and loan credit commitments, the lower limits for the credit they are referred to is the same as the one for the credits reported to Centrale dei Rischi.

For the interest rates on deposits, the observations only refer to the deposits with amount larger than 20 millions Italian lira.

Loan cash credits until December 31st 1996 include overdue payments, while from from January 1997 the latter are reported separately. This generates another time discontinuity for all the loan sizes. data for net interbank deposits and foreign deposits include interbank transfers, while any other variable in the dataset does not include them.

C) SAMPLING AND CLASS CRITERIA FOR THE VARIABLES OF THE DATASE

All the size classes have been calculated on the basis of loan cash credits granted by the whole

banking system to each customer (see tables below).

For bad debts, since 1997 the distribution of the borrowers among different classes is determined on the basis of the global amounts for which each borrower has overdue payments.

For what concerns loan credit commitments, since 1997 they have no longer been distributed (in the survey form the Bank of Italy) by class size of the credit lines. For this reason, since 1997 the only observation available from the Bank of Italy was the total credit to all of the borrowers, and this specific variable could not be used for the estimates. Therefore the data employed in the econometric analyses are only based on the loan cash credits.

The geographic areas (see tables below) refer to the localization of the local agency of the bank reporting the data to the statistical survey.

The distribution by size classes for banks (see tables below) has been updated in 1995 and also includes the former "long term credit special institutions" and "special credit sections" (see "Supplementi al Bollettino Statistico", Banca d'Italia, Anno V Numero 32 - 16 June 1995). For the previous periods the size classification only concerned ordinary banks, which broadly correspond to the present "banks with short run deposits".

We report below a few tables summarizing all the criteria of the statistical surveys form the Bank of Italy and all the definitions concerning the size classes and geographic areas.

Because of the discontinuities of the definitions in the Bank of Italy dataset, the Bank of Italy size classes "1" and "2" have been aggregated into the size class "c1" for the analyses performed here. This means that the corresponding interest rate is the weighted average of the interest rates of the two corresponding classes. A similar aggregation has been done for the bank deposits and the corresponding interest rates, which, in any case, turned out to be always non significant in preliminary econometric analyses referred to the interest rate on bank credit and to the overdrafts granted to the largest size class of customers. Furthermore, since the interest rate on deposits are (in Italy) very stable and show very little variability, they have not been the object of specific descriptive analyses.

Size classes of loan credits (DEFINITIONS OF THE BANK OF ITALY)	Description of the size classes employed by the Bank of Italy	Observed variable	Definition for the corresponding classes of loans included in the tables and graphics and for the dummy variable included in the estimates
1	Less than 80 million lira (between 1990 and 1995); less than 150 million lira (between 1996 and 1998)	outstanding amount of loan credit	c1 (in our tables the class c1 includes the sum of the Bank of Italy size classes "1" and "2")
2	From 80 to 250 millions lira (until 1995); from 150 to 250 million lira (between 1996 and 1998)	outstanding amount of loan credit	c1 (in our tables the class c1 includes the sum of the Bank of Italy size classes "1" and "2")
3	From 250 to 500 million lira	outstanding amount of loan credit	c3
4	From 500 million lira to 1 billion	outstanding amount of loan credit	c4
5	over 1 billion	outstanding amount of loan credit	c5

Size classes of loan credits to which the interest rates are referred (DEFINITIONS OF THE BANK OF ITALY)	Description of the size classes employed by the Bank of Italy	Observed variable	Definition for the corresponding classes of loan credits to which the interest rates refer in the tables and graphics
1	Less than 100 million lira (until 1996); less than 150 million lira (from 1997)	Average interest rate on loan credits of the given size class	c1
2	From 100 to 250 million lira (until 1996); from 150 to 250 million lira (from 1997)	Average interest rate on loan credits of the given size class	c1
3	From 250 to 500 million lira	Average interest rate on loan credits of the given size class	c3
4	From 500 million lira to 1 billion	Average interest rate on loan credits of the given size class	c4
5	over 1 billion	Average interest rate on loan credits of the given size class	c5

Size Classes of Banks	Description of the exact definition employed by the Bank of Italy	Definition for the corresponding size classes of banks included in the tables and for the dummy variable included in the estimates
1	major banks and large sized banks	I1
2	Average sized banks	I2
3	Small and minor sized banks	I3

Geographic area according to the	Regions included in the geographic area	Definition for the dummy variable included in the estimates
1	Piemonte, Val d'Aosta, Liguria	a1
2	Lombardia, Triveneto, Emilia Romagna	a2
3	Toscana, Marche, Umbria	a3
4	Lazio	a4
5	Abruzzi, Molise, Puglia, Basilicata	a5
6	Campania, Calabria, Sicilia, Sardegna	a6

Size classes of bank deposits the observed interest rates are referred to	Description of the size classes employed by the Bank of Italy	Observed variable
1	Less than 100 million lira	Average interest rate on bank deposits of the given size class
2	From 100 to 250 million lira	Average interest rate on bank deposits of the given size class
3	From 250 to 500 million lira	Average interest rate on bank deposits of the given size class
4	From 500 million lira to 1 billion	Average interest rate on bank deposits of the given size class
5	Over 1 billion	Average interest rate on bank deposits of the given size class

D) A BRIEF SUMMARY OF DESCRIPTIVE STATISTICS

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN ALL THE GEOGRAPHIC AREAS

	loans from l1 to c1	loans from l2 to c1	loans from l3 to c1
average interest rate for c1 for all the areas	16.86705	16.52475	15.91545
variance for c1 for all the areas	2.541764	2.170026	3.577675
	loans from l1 to c3	loans from l2 to c3	loans from l3 to c3
average interest rate for c3 for all the areas	15.60936	15.15257	14.89092
variance for c3 for all the areas	4.048736	3.552005	4.42039
	loans from l1 to c4	loans from l2 to c4	loans from l3 to c4
average interest rate for c4 for all the areas	15.16653	14.72696	14.27933
variance for c4 in all the areas	3.9862	3.599013	4.3118
	loans from l1 to c5	loans from l2 to c5	loans from l3 to c5
average interest rate for c5 in all the areas	12.85701	12.54965	12.55033
variance for 5 in all the areas	4.734899	4.039671	4.500353

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 1

	all bank size	from l1	from l2	from l3
average interest rate for c1 in a1	15.7431	15.95678	15.72506	15.54746
var. for c1 in a1	1.979536	2.012298	2.008203	2.002315
average interest rate for c3 in a1	14.40492	14.78237	14.26627	14.16611
var. for c3 in a1	3.162326	3.204932	3.63162	2.990966
average interest rate for c4 in a1	13.93973	14.31103	13.84119	13.66697
var. for c4 in a1	3.288572	3.181843	3.54873	3.111772
average interest rate for c5 in a1	11.93172	11.99015	12.01479	11.79021
var. for c5 in a1	3.848068	4.354332	3.788413	3.610673

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 2

	all bank size	from l1	from l2	from l3
average interest rate for c1 in a2	15.77032	15.50911	15.98415	15.81769
var. for c1 in a2	2.022209	2.016239	2.150746	1.916452
average interest rate for c3 in a2	14.50204	14.37267	14.71491	14.41854
var. for c3 in a2	3.301466	3.298332	3.07674	3.664488
average interest rate for c4 in a2	13.87443	14.11513	14.14424	13.95566
var. for c4 in a2	3.426298	3.450414	3.294419	3.726946
average interest rate for c5 in a2	11.94183	11.83155	11.92504	12.0689
var. for c5 in a2	4.147299	4.731783	3.950131	3.989705

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 3

	all bank size	from l1	from l2	from l3
average interest rate for c1 in a3	16.26506	16.62611	16.54421	15.62488
var. for c1 in a3	2.341941	1.759564	2.220764	2.604152
average interest rate for c3 in a3	14.86557	15.31585	14.95864	14.32222
var. for c3 in a3	3.951888	3.490071	4.638095	3.452015
average interest rate for c4 in a3	14.4187	14.83658	14.56143	13.85808
var. for c4 in a3	4.103046	3.676795	4.779664	3.583915
average interest rate for c5 in a3	12.38576	12.8429	12.37386	11.94053
var. for c5 in a3	4.419008	4.480293	4.474449	4.158398

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 4

	all bank size	from I1	from I2	from I3
average interest rate for c1 in a4	16.792	17.39481	16.87112	16.11008
var. for c1 in a4	1.747248	1.325765	1.652238	1.544723
average interest rate for c3 in a4	15.52946	15.9933	15.83293	14.76216
var. for c3 in a4	3.169065	3.361136	2.390434	3.029693
average interest rate for c4 in a4	15.04198	15.52703	15.2425	14.3564
var. for c4 in a4	3.163659	3.426189	2.334444	3.159273
average interest rate for c5 in a4	12.38232	12.55527	12.26988	12.32182
var. for c5 in a4	3.458965	4.094882	3.051656	3.398884

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 5

	all bank size	From I1	from I2	from I3
average interest rate for c1 in a5	17.86528	17.79964	17.24831	18.5479
var. for c1 in a5	2.558648	1.794134	1.957937	3.258426
average interest rate for c3 in a5	16.21687	16.47638	15.74477	16.42947
var. for c3 in a5	4.609047	3.741485	3.652608	6.375212
average interest rate for c4 in a5	15.63421	16.04618	15.51343	15.34301
var. for c4 in a5	4.753072	3.643787	3.088822	7.546154
average interest rate for c5 in a5	13.61227	13.98716	13.1621	13.68754
var. for c5 in a5	4.242804	3.608241	3.388026	5.637557

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

MEAN AND VARIANCE OF THE INTEREST RATES SET ON AVERAGE BY THE BANKS OF VARIOUS SIZE FOR THE VARIOUS SIZE CLASSES OF LOAN CREDITS IN AREA 6

	all bank size	from I1	from I2	from I3
average interest rate for c1 in a6	17.04591	17.91583	16.77564	16.44625
var. for c1 in a6	2.309148	1.699835	1.792189	2.395098
average interest rate for c3 in a6	15.78683	16.71556	15.39789	15.24702
var. for c3 in a6	3.761436	3.68598	2.898527	4.006274
average interest rate for c4 in a6	15.23935	16.16324	15.05896	14.49585
var. for c4 in a6	3.62321	3.197539	2.897184	3.517479
average interest rate for c5 in a6	13.66008	13.93501	13.55223	13.49301
var. for c5 in a6	3.55933	3.410671	3.983223	3.387823

N.B. mean and variance referred to the size class "c1" cannot be compared with the other because it refers to a different and shorter observation period, where the interest rates have been, on average, higher.

GENERAL UNRESTRICTED MODEL
EQUATION FOR THE SPREAD BETWEEN INTEREST RATES ON BANK OVERDRAFT AND
MONETARY POLICY RATE (INTEREST RATE ON LOANS BY THE BANK OF ITALY)
The data employed for the estimates cover the period 1990 QI - 1998 QIV

D.P.D. RESULTS

LEVELS IV

Number of firms: 54 Sample period is 1991 QII to 1998 QI
Observations: 1512 Degrees of freedom: 1487

Dependent variable is: ta_d

Instruments used are:

CONST ta_d(-2) ta_d(-3) ta_d(-4) ta_d(-5) so_u(-1) so_u(-2) so_u(-3)
so_u(-4) so_u(-5) isco(-1) tdif(-1) tdif(-2) tdif(-3) tdif(-4) tdif(-5)
c3 c4 a1 a3 a4 a5 a6 IND DUMS

ONE-STEP ESTIMATES WITH ROBUST TEST STATISTICS

Wald test of joint significance: 505.422951 df = 22
Wald test - jt sig of ind dums: 0.341599 df = 2
Wald test selected by user: 10.299139 df = 10
Testing: ta_d(-2) ta_d(-3) ta_d(-4) so_u(-2) so_u(-3) so_u(-4) isco
tdif tdif(-1) tdif(-2)

Var	Coef	Std. Error	T-Stat	P-Value
CONST	0.918082	5.373138	0.170865	0.864330
ta_d(-1)	3.237775	3.886986	0.832978	0.404857
ta_d(-2)	-0.970432	1.540478	-0.629955	0.528724
ta_d(-3)	-0.157001	0.433125	-0.362485	0.716989
ta_d(-4)	-0.396844	0.909495	-0.436335	0.662594
so_u	5.501375	3.635302	1.513320	0.130198
so_u(-1)	-3.951405	3.237292	-1.220590	0.222241
so_u(-2)	-1.606319	3.287958	-0.488546	0.625163
so_u(-3)	1.047737	2.816219	0.372037	0.709866
so_u(-4)	-3.611479	3.118469	-1.158093	0.246826
isco	0.046276	0.125520	0.368677	0.712369
tdif	-2.573195	3.468352	-0.741907	0.458144
tdif(-1)	0.662969	2.671158	0.248196	0.803983
tdif(-2)	0.474065	0.452225	1.048293	0.294504
tdif(-3)	0.283370	0.289738	0.978022	0.328063
tdif(-4)	0.339537	0.287402	1.181402	0.237443
c3	-3.979281	2.735695	-1.454577	0.145786
c4	-3.158094	2.214824	-1.425889	0.153900
a1	0.022598	0.234861	0.096217	0.923349
a3	-0.057324	0.479939	-0.119441	0.904926
a4	-0.255889	0.967343	-0.264527	0.791374
a5	-0.694210	1.738182	-0.399389	0.689607
a6	-0.911166	1.378700	-0.660888	0.508684
I2	-0.140736	0.608300	-0.231359	0.817036
I3	-0.416781	1.109827	-0.375536	0.707262

Robust test for first-order serial correlation: -0.814 [54]
Robust test for second-order serial correlation: 0.620 [54]

GENERAL UNRESTRICTED MODEL
EQUATION FOR THE SHARE OF CREDIT GRANTED TO THE LARGEST SIZE CLASS
("CLIENTS 5")

The data employed for the estimates cover the period 1990 QI - 1998 QIV

D.P.D. RESULTS

LEVELS IV

Number of firms: 18 Sample period is 1991 QII to 1998 QI
Observations: 504 Degrees of freedom: 473

Dependent variable is: c5_t

Instruments used are:

CONST c5_t(-2) c5_t(-3) c5_t(-4) c5_t(-5) tat(-1) tat(-3) tat(-4)
tat(-2) tat(-5) so_u(-1) so_u(-2) so_u(-3) so_u(-4) so_u(-5) tan(-1)
tan(-2) tan(-3) tan(-4) tan(-5) isco(-1) tma(-1) tma(-2) tma(-3)
tma(-4) tma(-5) a1 a3 a4 a5 a6

ONE-STEP ESTIMATES WITH ROBUST TEST STATISTICS

Wald test of joint significance: 4431.770670 df = 30
Wald test selected by user: 7.462959 df = 24
Testing: c5_t(-2) c5_t(-3) c5_t(-4) tat tat(-1) tat(-2) tat(-4)
so_u so_u(-1) tan tan(-1) tan(-2) tan(-3) tan(-4) tma
tma(-1) tma(-2) tma(-3) tma(-4) a1 a3 a4 a5
a6

Var	Coef	Std. Error	T-Stat	P-Value
CONST	-0.138288	0.156984	-0.880902	0.378371
c5_t(-1)	2.942707	2.112914	1.392725	0.163703
c5_t(-2)	-2.005299	2.092366	-0.958388	0.337867
c5_t(-3)	0.174508	0.296365	0.588829	0.555976
c5_t(-4)	-0.011989	0.141317	-0.084839	0.932390
tat	0.027344	0.027825	0.982712	0.325749
tat(-1)	-0.013846	0.014940	-0.926768	0.354047
tat(-2)	0.004455	0.010317	0.431821	0.665872
tat(-3)	-0.003635	0.005215	-0.697092	0.485745
tat(-4)	-0.003304	0.006173	-0.535252	0.592476
so_u	-0.081250	0.166762	-0.487221	0.626102
so_u(-1)	0.120028	0.180136	0.666317	0.505209
so_u(-2)	-0.060010	0.052729	-1.138082	0.255086
so_u(-3)	0.084269	0.064958	1.297276	0.194536
so_u(-4)	-0.119746	0.099812	-1.199712	0.230251
tan	-0.007168	0.014392	-0.498050	0.618448
tan(-1)	0.008115	0.013108	0.619095	0.535854
tan(-2)	0.001665	0.004124	0.403840	0.686330
tan(-3)	0.000392	0.001968	0.199169	0.842131
tan(-4)	0.001837	0.002366	0.776374	0.437528
isco	0.001490	0.001720	0.866275	0.386339
tma	-0.024440	0.023699	-1.031243	0.302427
tma(-1)	0.007476	0.009783	0.764128	0.444791
tma(-2)	0.002505	0.006351	0.394353	0.693320
tma(-3)	-0.001077	0.003461	-0.311090	0.755733
tma(-4)	0.005157	0.006172	0.835599	0.403381
a1	0.000255	0.003423	0.074583	0.940547
a3	0.006086	0.007869	0.773465	0.439247
a4	0.003671	0.004724	0.777195	0.437044
a5	0.010938	0.013995	0.781585	0.434458
a6	0.002276	0.009911	0.229597	0.818405

Robust test for first-order serial correlation: -0.978 [18]
Robust test for second-order serial correlation: 1.005 [18]

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